



The Structure of the Chinese Material Value Scale: An Eastern Cultural View

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We contend that Chinese materialism will demonstrate a unique structure owing to its culture (collectivism). Chinese culture has affected people's attitudes toward money, possession, success, and happiness for over a 1000 years (Bauer, 1976), which suggests a uniquely Chinese lifestyle different from that of Western culture. We therefore propose that the structure of materialism in China is different from that in Western countries. The present study was designed to investigate the dimensionality of materialism in Chinese samples to uncover culture-specific structural features.

CONCEPTUAL BACKGROUND

Materialism and Material Value

Scholars have historically found that materialism is rooted in the West, where the modern industry and contemporary happiness-seeking lifestyle originated (Belk, 1985). It is consistently agreed that materialism is a product of economic development and market maturity. According to Engel's coefficient, the per-capita GDP of \$3,000 is an important economic development, signifying that the economy has reached the level of middle-income countries and that economic and social life in China will change. When referring to consumer behavior, people pay more attention to quality of life and consumption diversification.

Many researchers have defined and measured materialism from the perspective of attitude, personality, and values. For example, Campbell (1970) developed a materialism scale which focused on attitudes. Moschis and Churchill (1978) developed a materialism scale that included attitudes toward money as well as attitudes toward possessions. Focusing on personality traits, Belk (1985) developed a 24-item materialism scale with three-dimensions—envy, possessiveness, and non-generosity—a structure that seems inconsistent with the nature of materialism in less hedonic societies. Later, Ger and Belk (1990)

1985), it is necessary to study the structure of MVS in Eastern cultures such as China.

Chinese Material Value

China is a society that emphasizes collectivism and personal social value (Hu, 1944; Ho, 1976; Hofstede, 1980; Krebs, 1982). Social evaluation is one of most important standards; people tend to feel success and happiness when they receive high levels of social evaluation. Possession, which can bring social status and reputation, is commonly used to evaluate success or happiness.

Chinese personal welfare, interest, and achievement are set collectively (Triandis, 1994); therefore, social status, roles, and relationships are explicit and public and play an important part in social life. In this context, social value is an important index of happiness and life satisfaction (Barger et al., 2009). Anything that can bring social status and reputation will bring happiness. Material success is one such thing. In conventional Chinese society, pursuing material success and surpassing others in life and career are the highest forms of success. The people with the highest wealth and status often represent the ultimate goal of happiness.

Economic conditions are the premise of materialism. In 2008, the Chinese per capita GDP entered the US \$3,000 range for the first time (2009 China National Bureau of Statistics). The consumption of the Chinese people is characterized by diversification, and the materialistic values are becoming more and more mature (Sun et al., 2014).

Based the above review of the literature, we investigated the hypothesis that for Chinese people, the success domain and happiness domain of the MVS will be more closely related to each other, especially when the success and happiness can be represented through public possessions. Given the likely overlap between the two domains, we speculate that in China, the structure of material value may be two-dimensional instead of three-dimensional.

OVERVIEW OF THE PRESENT RESEARCH

We conducted two studies in mainland China. Study 1 examined the structure of Richins and Dawson's (1992) MVS and its short version (Richins, 2004) in Chinese samples and revealed that there were only two-dimensions of materialism in China with the finding that the happiness factor and success factor merged together. In Study 2, we examined how strongly the items in the happiness and success domains of the MVS were associated with the concept of success and happiness through an implicit experiment, the dot-probe task. This further verified the overlapping of the items from the happiness and success subscales of the MVS.

The studies were approved by the Ethics Committee of Peking University and Tsinghua University. In Study 1, data were collected from different cities in Mainland, for example, Guangzhou, Beijing, and Shenzhen. Before completing the questionnaire, all participants read informed content and knew their rights. In Study 2, the participants were recruited from a

public university. All participants gave written informed consent. The participants were informed that the questionnaire and the experiment did not involve any privacy. The participants could withdraw or terminate the experiment when they experience any discomfort during the experiment.

STUDY 1: THE APPLICABILITY OF RICHINS AND DAWSON'S MVS IN CHINESE SAMPLES

The aim of Study 1 was to examine the dimensionality of Richins and Dawson's MVS in Chinese samples. We argue that Chinese people may have a unique structure of materialism, and the three-component structure of Richins and Dawson's MVS may not fit a Chinese sample.

Method

Participants

A total of 981 consumers from different cities in mainland China participated in the study. Participants included employees ($N = 857$), and students ($N = 124$). The mean age for the sample was 29.43 ($SD = 7.73$) and ranged from 18 to 78 years old. Among these participants, 452 were male, 68.4% had special training experience, and 32.2% had bachelor's or higher degrees.

The participants were divided into two samples according to the number of questionnaire, namely, even-numbered sample ($N = 490$) and uneven-numbered sample ($N = 491$). One sample's data was used for exploratory factor analysis, and the other sample's data were used for confirmatory factor analysis (CFA).

In this study, the skewness coefficient of each item is between 0.04 and 0.72, and the absolute value of the Kurtosis is between 0.51 and 1.12. According to West et al. (1995), when the skewness and Kurtosis are less than 2 and 7, maximum likelihood estimation is robust. Thus, maximum likelihood estimation was used to handle missing data.

Measures

Material value scale

Material value scale (MVS) was developed by Richins and Dawson (1992). In 2004, Richins developed a 15-item short version.

In order to establish translation equivalence, Mullen's (1995) method was used to translate the 18 original items of the Richins and Dawson's (1992) scale into Chinese. A five-point Likert scale was used for all item ratings (1 = *strongly agree*, 5 = *strongly disagree*).

After all participants completed the MVS, the following scales were also tested to examine the validity of MVS. Of our participants, 375 completed the Oxford Happiness Questionnaire (OHQ) and Money Attitude Scale (MAS), and 145 completed the Money Ethic Scale (MES) and Social Comparison Scale.

Oxford happiness questionnaire

Psychological well-being was assessed with the OHQ (Hills and Argyle, 2002), devised as a broad measure of personal happiness.

The Chinese version (Shi et al., 2005; Li and Chen, 2013) was used in the present study; the alpha coefficient was 0.90. Previous studies have revealed that individuals high on materialism have low psychological well-being (Richins and Dawson, 1992; Burroughs and Rindfleisch, 2002; Dittmar et al., 2014; Nagpaal and Pang, 2017).

Money attitude scale

The MAS was developed by Yamauchi and Templer (1982). The MAS has 29 items rated on a seven-point Likert scale ranging from “always” to “never.” This concept is indirectly related to materialism (Belk, 1985). The MAS has two subscales: Power–Prestige (measuring the degree to which one sees possession as a tool to influence and impress others) and Retention–Time (measuring the degree to which one plans for their financial future). We calculated both the total-scale and subscale scores to examine the criterion-related validity of the MAS. The alpha coefficient for total scale was 0.87; for the two subscales, the alpha coefficients were 0.89 and 0.82, respectively. The Chinese version (Hong et al., 2009) was used in the present study.

Money ethic scale (MES)

Tang (1993, 1995) developed the MES according to the ABC model of attitude toward money, which includes affective, behavioral, and cognitive components. The Chinese version (Wang, 2005) was used in the present study. We used the 15-item, five-factor (budget, evil, equity, success, and motivation) scale revised by Wang (2005), rated on a five-point Likert scale (Tang et al., 2003). The alpha coefficient for the total scale was 0.85. We predicted that the MES would be correlated with the MVS—the more people value money, the more inclined they are to pursue material wealth and define happiness and success by money, and the more they value material items.

Social comparison scale

Gibbons and Buunk (1999) developed the Social Comparison Scale, which includes 11 items. The Chinese version revised by Wang et al. (2006) was used in this study. People with high levels of social comparison are sensitive to the behavior of others, have a degree of uncertainty about themselves, and try their best to reduce self-uncertainty. They are interested in how others perform. The alpha coefficient for the total Social Comparison Scale was 0.85. With Westerners, East Asians differed in satisfaction with wealth regardless of social comparative experience (Ogden and Venkat, 2001). Status consumption is positively correlated to MVS and social comparison, and social comparison reflects the consumer behavior of pursuing status (Heaney et al., 2005).

Procedure

The MVS was distributed to the participants along with a consent form and demographic information sheet. Then other questionnaires for validity testing were completed. All participants were informed that completing the questionnaire was voluntary and anonymous.

Results

Confirmatory Factor Analysis of the MVS

First, we examined Richins and Dawson's (1992) MVS in Sample 1. The internal consistency (Cronbach's alpha) was 0.80 for the 18-item scale, with 0.63, 0.63, and 0.54 for the three subscales, respectively; for the 15-item scale, Cronbach's alpha was 0.79 for the overall scale and 0.62, 0.52, and 0.54 for the three subscales, respectively. Comparing with previous studies, the internal reliability of Richins and Dawson's original MVS is better than this one. The results of the CFA of Richins and Dawson's MVS, performed with Lisrel 8.80, revealed an unsatisfactory factor structure, with all absolute and incremental index values failing to meet the desirable levels. Specifically, the χ^2 was highly significant ($\chi^2 \geq 829.06$, $df \geq 132$, $p < 0.001$), and the χ^2/df value ≥ 6.28 , which was higher than the desired value of 3.0. (Wheaton, 1987). Moreover, the RMSEA was 0.103, indicating that the model of fit cannot be accepted; the GFI was 0.84. In addition, the three incremental fit indices (CFI ≥ 0.79 , NNFI/TLI ≥ 0.75 , RMSEA > 0.1) failed to reach the recommended criterion of 0.90. Additionally, the three subscales were highly correlated with each other ($r_{12} \geq 0.97$, $r_{13} \geq 0.87$, $r_{23} \geq 0.84$). We also conducted CFA for Richins' (2004) short-version MVS. The results, similar to those of the long version, also revealed an unsatisfactory factor structure (see **Table 1**).

The results demonstrate that the Richins and Dawson's (1992) MVS may have a different structure in Chinese samples than in American samples. The principal-component analysis was conducted with oblique rotation to identify the MVS dimensionality in Chinese samples. No restrictions were set limiting the number of factors in the analysis; thus, the items were free to load on any number of factors. The criterion for item inclusion was set with factor loadings > 0.30 . Principal component analysis showed that there are two-factors whose eigenvalues were greater than one and those were factors were rotated by the method of Varimax with Kaiser Normalization.

The results of EFA yielded a two-factor solution with the 18 or 15 items grouping into two coherent and interpretable factors: (1) Happiness–Success (10 items in the 18-item version and 9 items in the 15-item version) and (2) Material Possession (8 items in the 18-item version and 6 items in the 15-item version). The two-factors accounted for 33.8% of the response variance in the 18-item version and 36.9% in the 15-item version.

The correlations of two subscale were low but statistically significant for both the 15-item ($r \geq 0.28$, $p < 0.001$, $df \geq 488$) and the 18-item inter-factor ($r \geq 0.32$, $p < 0.001$, $df \geq 488$). According to Devellis (1991), the subscales demonstrated acceptable levels of internal consistency (Cronbach's alpha) with Happiness–Success ($\alpha \geq 0.80$ for the 18-item and $\alpha \geq 0.79$ for the 15-item) and Material Possession ($\alpha \geq 0.67$ for the 18-item and $\alpha \geq 0.62$ for the 15-item). For each version, all items loaded on the two-factors consistently and stably, without any double loading.

Compared with the three-factor model, the two-factor model was confirmed to be more suitable. This indicates that materialism values have a simpler structure in Chinese samples than in American samples. Specifically, the happiness and success domains converged as one-factor. To reexamine the new two-factor structure, we conducted a confirmative factor analysis

TABLE 1 | Indicators of standardized CFA.

Original MVS (<i>N</i> = 490, <i>p</i> < 0.001)						
Model	Items	χ^2/df	GFI	CFI	NNFI	RMSEA (90%, CI)
Three-factor	18	948.19/132	0.83	0.76	0.75	0.103 (0.10, 0.11)
	15	491.93/87	0.88	0.84	0.80	0.10 (0.09, 0.10)
Revised MVS(<i>N</i> = 491, <i>p</i> < 0.001)						
Model	Items	χ^2/df	GFI	CFI	NNFI	RMSEA (90%, CI)
Two-factor	18	553.05/134	0.89	0.87	0.85	0.08 (0.07, 0.09)
	15	325.22/89	0.92	0.91	0.90	0.07 (0.06, 0.08)

χ^2 , normal theory weighted least squares chi-square; *df*, degrees of freedom; ($\chi^2/df < 3$); GFI, goodness-of-fit index (0.90 and above); CFI, comparative fit index (0.90 and above); NNFI, Non-Normed Fit Index (0.90 and above); RMSEA, root mean square error of approximation (RMSEA < 0.08).

using the data from Sample 2. The results showed that the two-factor model for both the long and short version, and the materialism value of the original and short-form questionnaire fit well (Table 1). Specially, the 15-item MVS had higher GFI (0.92), critical CFI (0.91) and NNFI (0.90), lower RMSEA (RMSEA D 0.07, 90% CI: 0.06–0.08) and inter-factor correlation (r_{12} D 0.45, $p < 0.01$). The results of CFA showed in Table 1. Table 2 lists the results for the two-factor structure using exploratory factors analysis. Combine multiple indicators of EFA and CFA, such as factor loading, RMSEA, CFI, NNFI (TLI) etc., the structure model was acceptable (Baumgartner and Homburg, 1996). Briefly, EFA and CFA supported the hypothesis that the 15-item short-form MVS had a two-factor structure in Chinese samples. This indicates that the Chinese MVS is different from that of Westerners.

Moreover, we investigated the reliability and validity of the two-factor Chinese MVS with 15 items. A total of 375 participants completed the OHQ and MAS, and 145 participants completed the MES and Social Comparison Scale.

Reliability

The alpha coefficient of the Chinese 15-item MVS was 0.81, and the alpha coefficient of the adjusted Spearman–Brown formula was 0.86. The alpha coefficient was 0.79 and 0.72 for the two-factors, respectively. After one and a half months, we randomly selected 52 participants to take the MVS again, and the test–retest coefficient was 0.87 (95%CI: 0.68–0.93, $p < 0.01$, *df* D 50).

Validity

Convergent validity

The Chinese MVS scores correlated with the OHQ scores (r D 0.13, $p < 0.05$, *df* D 373). The two subscale scores were also correlated with OHQ (H&S: r D 0.17, $p < 0.05$, *df* D 373).

The Chinese MVS scores correlated with the MAS scores (r D 0.15, $p < 0.05$, *df* D 373).

The Chinese MVS scores correlated with the MES scores (r D 0.15, $p < 0.05$, *df* D 373).

The Chinese MVS scores correlated with the Social Comparison Scale scores (r D 0.15, $p < 0.05$, *df* D 373).

Explored (value of χ^2/df) MES (227 *df* = 200 χ^2/df = 1.135 p = 0.001) MES (227 *df* = 200 χ^2/df = 1.135 p = 0.001) MES (227 *df* = 200 χ^2/df = 1.135 p = 0.001) MES (227 *df* = 200 χ^2/df = 1.135 p = 0.001)

The Cse 750 *df* = 180 χ^2/df = 4.167 p = 0.001) MES (227 *df* = 200 χ^2/df = 1.135 p = 0.001) MES (227 *df* = 200 χ^2/df = 1.135 p = 0.001) MES (227 *df* = 200 χ^2/df = 1.135 p = 0.001)

represents long-term orientation. The Chinese MVS scores were also significantly correlated with MES scores ($r = 0.43, p < 0.01, df = 143$). Subscale 1 (H&S) was positively correlated with MES ($r = 0.49, p < 0.01, df = 143$), while Subscale 2 (MP) was insignificantly negatively correlated with MES. The Chinese MVS were also positively correlated with Social Comparison Scale scores ($r = 0.38, p < 0.01, df = 143$). These results showed that the two-factor structure Chinese version of the MVS had acceptable reliability and validity. The correlation coefficients and confidence interval of criterion validity were list in Table 3 as following.

Discussion

The results of Study 1 showed that happiness and success converged as one-factor in the Chinese sample while possession centrality remained an independent factor. This was inconsistent with previous findings in American studies that the MVS had three-factors, each with a clearly separate construction (e.g., Richins and Dawson, 1992; Richins, 2004). This inconsistency can be explained from three perspectives. First, the meanings of happiness and success and their subjective and objective expressions correlate to each other in some ways. Much research has indicated that the achievement of a successful career and personal goals is an important component of individual subjective well-being (e.g., Campbell et al., 1976; Lu and Shih, 1997; Furnham and Cheng, 2000a), that material possession plays an important role in happiness and success (e.g., Lu and Shih, 1997; Furnham and Cheng, 2000b), and that the individual's achievement and material wealth exemplify not only happiness but also success (Argyle et al., 1989; Chiasson et al., 1996; Fan and Karnilowicz, 1997; Furnham and Cheng, 2000b; Lu, 2001).

With respect to materialism values, people enjoy material wealth as the foundation of happiness, and the objective expression of happiness is the foundation of psychological well-being. Moreover, the success that material wealth brings is an important component of an individual's subjective well-being (Argyle et al., 1989; Chiasson et al., 1996; Fan and Karnilowicz, 1997; Furnham and Cheng, 2000a; Lu, 2001), objectively expressed as happiness. Therefore, the success and happiness that material wealth brings may relate to each other by both contributing to people's psychological well-being.

The second explanation comes from the economical perspective. It has been shown that Teenagers growing up with low socioeconomic status are inclined to regard material wealth as their most important goal (Singhal and Misra, 1992; Kasser et al., 1995). In rapidly developing societies, people crave material wealth, especially in countries such as China that experience sharply accelerating urbanization and commercialization (Ger and Belk, 1996). This, to a certain extent, also promotes the development of Chinese materialism values and strengthens the idea that people use material wealth as one of the important standards of success and happiness.

Third, from a cultural perspective, Westerners regard their own happiness as an ultimate life goal. The individual is independent, having unique inherent characteristics—such as

personality, ability, motive, and values—and his/her behaviors stem from the influence of these inherent characteristics (Sampson, 1988). Comparatively, East Asians consider whether their views or behaviors are in agreement with their in-groups or societies (Triandis et al., 1990); their happiness is influenced by external standards. For instance, societal approval, maintenance of public image, modesty, and respect of parents and elders are all considered to be important values. Therefore, material wealth that is apparent to the public is apt to become an important factor of Chinese life goals that connect success and happiness.

In sum, the meanings of happiness and success are similarly expressed in China. Based on scales measuring material values, happiness and success were not independent; instead, they merged together as one-factor. The question of how the two-dimensions merged remains, however. Do they simply overlap, or does one comprise the other? Does this phenomenon occur due to the specific items? These questions need to be examined by empirical studies.

STUDY 2: VERIFYING THE DIMENSION MERGENCE VIA IMPLICIT STUDY

In Study 1, we found that the Chinese MVS had only two-factors, with happiness centrality and success centrality converged into one-factor. However, Study 1 used an explicit self-report method. Some research indicates that self-report measurements can be influenced by social desirability (Himmelfarb and Lickteig, 1982; Paulhus, 1984; Schlenker and Weingold, 1989). In order to confirm whether Chinese people really think the concepts of happiness and success are similar, we conducted an implicit experiment. We adopted the dot-probe paradigm (Macleod et al., 1986) used to investigate attention toward emotions, such as anxiety symptomatology, depression, social threat, and positive words in attention bias studies (e.g., Macleod et al., 1986; Gotlib et al., 1988; Ruiz-Caballero and Bermudez, 1997; Bradley et al., 1998; Mogg and Bradley, 1999). Some studies have used this paradigm to test whether subjects show an attentional orientation

TABLE 3 | The correlation between MVS and criterion validity.

		<i>N</i> = 375				<i>N</i> = 145	
		OHQ	MAS	Power-Prestige	Retention-Time	MES	SCS
MVS	Pearson <i>R</i>	0.13	0.32	0.45	0.02	0.43	0.38
	95% CI upper	0.22	0.22	0.38	0.02	0.30	0.21
	95% CI lower	0.01	0.40	0.52	0.19	0.51	0.48
MVS- H&S	Pearson <i>R</i>	0.17	0.46	0.56	0.10†	0.49	0.33
	95% CI upper	0.26	0.38	0.49	0.22	0.38	0.20
	95% CI lower	0.04	0.55	0.62	0.01	0.57	0.44
MVS- MP	Pearson <i>R</i>	0.10	0.02	0.11	0.11†	0.06	0.20
	95% CI upper	0.21	0.09	0.01	0.08	0.01	0.09
	95% CI lower	0.02	0.11	0.21	0.13	0.13	0.31

$p < 0.05$; $p < 0.01$; † $p < 0.10$.

would be shorter. If the participants focus on the meaning of an item in terms of both happiness and success, there is no attention bias toward the left or the right and no significant difference in RT. Based on this experimental paradigm, we examined whether and to what extent each item was implicitly associated with happiness and success. This allowed us to uncover how the success and happiness factors in the MVS relate to each other.

Method

Participants

Thirty-two undergraduates enrolled at a large public university (14 men, 18 women) with an average age of 20.29 ($SD = 1.51$, range 18–23 years) volunteered for Study 2. The participants who completed the experiment were paid for 10RMB. All participants were right-handed, had normal or corrected-to-normal visual acuity, and were naïve to the purpose of the experiment.

Apparatus and Materials

The modified probe detection task was programmed and presented using Presentation software (Schneider et al., 2002) at a 1,024 × 768 pixel screen. Participants were seated approximately 60 cm behind the computer screen.

The materials included priming sentences and category words. Ten priming sentences were taken from the two-dimensions of the Material Value Scale: happiness centrality and success centrality (e.g., own expensive cars, homes, and clothes). These sentences had been adapted to make all items appropriately equal or similar in length and in affirmativeness of the statement. Immediately after priming sentences, the category words “success” and “happiness” were presented in font size 30, 10 cm from the middle of the screen. Finally, a dot probe (up arrow or down arrow with 38 × 50 pixels) was presented, placed in the position of one of the category words. Participants were asked to respond by pressing either the up or down arrow button as quickly and accurately as possible, pressing the up arrow when the arrow on the screen pointed up and the down arrow when the arrow pointed down. Before each pair of category words

and arrow was presented, a fixed cross (0.6 cm × 0.6 cm) appeared in the center of screen to attract the participant’s gaze.

In the pilot study, 16 students judged the consistency between each priming sentence and its original item in the MVS in a matching task. They were asked to draw a line between one of the 10 edited priming sentences (for example, own expensive homes, cars, and clothes.) in one column and one of the 10 original items (e.g., I admire people who own expensive homes, cars, and clothes.) from the MVS in another column if they believed the two had similar meanings. The results showed that, on average, participants correctly matched 90% of the items. We then revised the priming sentence according to the results to make sure the shortened sentence used for priming retained the same meaning as its original item in the MVS.

Design and Procedure

Participants responded to six items during practice and 10 items in the formal experiment; each item was repeated 10 times. Priming sentences were presented by the computer one by one. The trials had no time limit; the sentence would not disappear from the screen until participants understood the meaning of the sentence and then pressed the space bar (see Figure 1). After refreshing, the screen showed a fixed cross (200 ms), followed by category words (800 ms), another fixed cross (200 ms), and an arrow, with the screen refreshing in-between. Each of the two category words (happiness and success) appeared on either the left or right side of the screen randomly; the arrow was placed in the position of any of the category words with equal probability and random order. The direction of the arrow was also random. Participants had to judge the direction of the arrow as quickly and accurately as possible. The entire experiment took around 15 min to complete.

We hypothesized that if participants thought the meaning of the item was success (happiness), their attention bias would be directed to the subsequent category word “success” (“happiness”), and the RT for judging the direction of the arrows would be faster when the arrow was presented in the position of the corresponding category word “success” (“happiness”) than that of “happiness” (“success”).

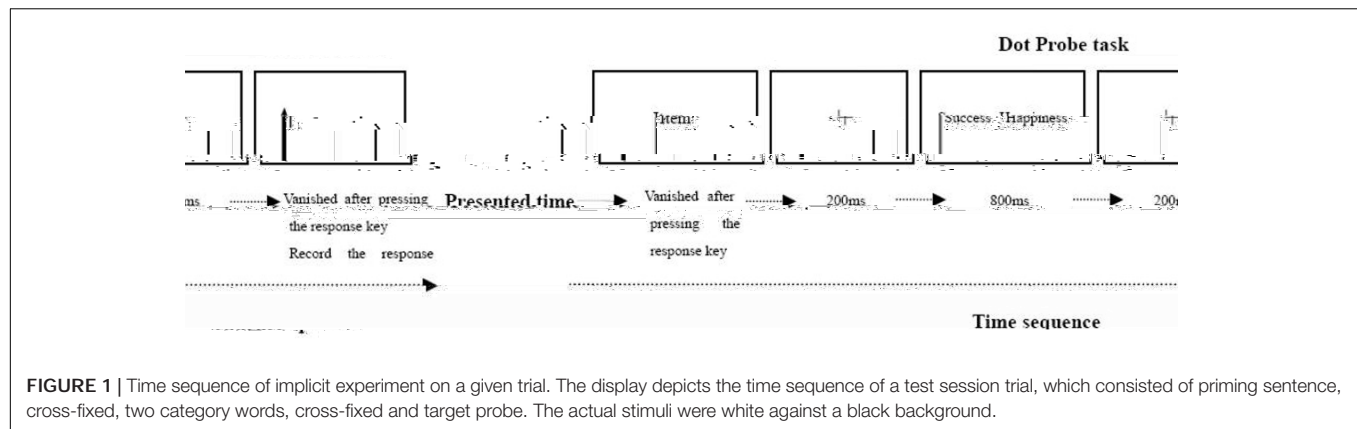


FIGURE 1 | Time sequence of implicit experiment on a given trial. The display depicts the time sequence of a test session trial, which consisted of priming sentence, cross-fixed, two category words, cross-fixed and target probe. The actual stimuli were white against a black background.

Results and Discussion

The error rate, RT, Mean, *SD*, and results of the *t*-test are listed in Table 4. The RTs were analyzed using a 2 (arrow position: consistent/inconsistent) × 2 (category word: success/happiness) repeated measures analysis of variance. Both variables were within subject factors. Erroneous responses were discarded from statistical analyses, as were trials with RTs shorter than 300 ms or longer than 2 standard deviations above each participant's mean RT.

The results showed a non-significant two-way interaction between arrow position and category word ($F[1,31] \leq 0.20$, $p > 0.05$, $\eta^2 \leq 0.01$). There was a significant main effect of category word ($F[1,31] \leq 4.57$, $p < 0.05$, $\eta^2 \leq 0.13$), and *ns* main effect for arrow position ($F[1,31] \leq 0.001$, $p > 0.05$, $\eta^2 \leq 0.001$).

We further analyzed the difference using a *t*-test, which showed that, for the success factor items of the MVS, there was no effect in the judgment of arrow direction ($t[31] \leq 0.42$, $p > 0.05$, Cohen's *d* ≤ 0.11). Participants showed no significant difference in response between the arrow placed in the position of category word "happiness" (494.39) and the arrow placed in the position of category word "success" (492.37). In other words, no attention bias occurred. This result indicates that participants equally related the success factor items of the MVS to the concepts of success and happiness.

However, for the happiness factor items of the MVS, there was an effect in the judgment of arrow direction ($t[31] \leq 2.09$, $p < 0.05$, Cohen's *d* ≤ 0.53); there was a significant difference in

participant response between the arrow placed in the position of category word "success" (494.57) and the arrow placed in the position of category word "happiness" (485.41). This result indicated that participants related the happiness items in the MVS more closely to the concept of happiness than to the concept of success.

These findings demonstrate that the Chinese views of happiness and success do relate to each other, confirming the result in Study 1 with implicit empirical evidence for convergence of material happiness and success. Moreover, this study shows that the relationship is not a simple overlaying between material happiness and success, but a comprising relationship. As the results show, material success is associated with the success and happiness concepts; that is, success is seen as indicating happiness, but happiness might not indicate success.

GENERAL DISCUSSION

The present research revealed that the MVS has a two-dimensional rather than three-dimensional structure in Chinese samples. Two of the original dimensions in Richins and Dawson's MVS, happiness centrality and success centrality, converged as one-factor. These findings were confirmed by two studies using both self-report questionnaire and controlled experiment.

In societies that attach importance to social evaluation criteria, success, and happiness tend to be integrated into one-dimension, so that the three-dimensions of materialistic value are reduced to 2. Analysis indicated that the 15-item Chinese MVS has better psychometric properties than 18-item Chinese MVS, original MVS and original short version MVS. This 15-item scale showed acceptable levels of reliability and validity, and good fitness of CFA. We verified the finding by using both explicit and implicit methods. Our finding reveals that Chinese materialism has unique characteristics, and may help shape the theory of material value into a more nuanced, culturally sensitive one.

Richins and Dawson's (1992) MVS has become one of the most widely used measurements of material value; it has shown good psychometric properties and has been validated in some cross-cultural studies. In the present study, the dimensionality and structure of the MVS were verified in China. We found no conflict

TABLE 4 | Response time results of implicit experiment.

		Happiness		Success	
		Mean	SD	Mean	SD
RT	Arrow directed success	494.57	58.34	492.37	56.10
	Arrow directed happiness	485.41	61.09	494.39	63.14
<i>t</i>		2.09		0.42	
<i>df</i>		31		31	
<i>p</i> -Value		0.028		0.68	
Error rate		11.2%		11.8%	

msec for RT. $p < 0.01$; $p < 0.05$.

with the previous material value theory developed by Richins and Dawson (1992). Moreover, we revealed new information that helps enrich the theory under the condition of social evaluation system. It is clear from our findings that cultural differences have an impact on the structure of materialism.

Moreover, the two-dimensional 15-item version has some advantages in application. First, the structure is clearer. It can explain the impact of culture on materialistic values. Secondly, it can save time and the capacity of the questionnaire. We call for more research to verify these advantages by comparing the two versions (original and short) in the future.

Materialism under Social Evaluation

Our observations showed that material possession centrality remains an independent factor under social evaluation, whereas the happiness and success dimensions may converge together. This suggests that the attribution of possession is culturally independent, while attributions of happiness and success are shaped by culture.

In Study 1, happiness and success items appeared in the same dimension. For example, “I’d be happier if I could afford to buy more things” represented happiness in the original MVS, and “Some of the most important achievements in life include acquiring material possessions” represented success in the original MVS; in our version, however, they represent the same factor.

The convergence of the material happiness and success factors lies in cultural influence. Easterners emphasize social role and public perspective when managing the self and hold interdependent self-concepts, which is directly related to the core feature of materialism: public consumption and possession (Marcus and Kitayama, 1991, 1994; Richins, 1994). This might be the cultural root through which people relate material success with material happiness. Success can bring people high social status, good reputation, and great wealth (Ting-Toomey, 1988, 1998). In Eastern culture, materialism is still portrayed as attention to status (Belk, 1985); for example, a large number of China’s advertisements represent the symbolic meaning of status. Good social status is a reasonable result of success and an acceptable demonstration of happiness. Furthermore, Easterners hold a holistic thinking style, which leads to a halo effect when we evaluate others by means of overall impression. Therefore, people tend to believe that possession success means happiness.

Methodological Implications for Future Research

We provided new methods to examine the structure of Richins and Dawson’s (1992) MVS in Chinese samples. Generally speaking, when measuring a concept using a scale, researchers have often used explicit Likert scales, which may be influenced by desirability and other biasing factors. This limits our understanding of the underlying logical relationships.

In the present research, we used both explicit and implicit experimental paradigms, which allowed us to double-check the overlapping relationship of factors in the MVS. This is particularly important in the research of values, since values are

relatively socially desirable, and participants’ explicit response may be more likely to be influenced by other motivations rather than their real personal views.

Limitation and Implications for Practice

This study explores the characteristics of material values of Chinese consumers from the perspective of culture. One contribution of this research is the recognition that consumers’ material values can be influenced by culture in economic life.

Several limitations are worthy of attention. The current study did not include samples from other countries, and there is also no comparison of cultural differences directly with other countries, the material values that explain multiculturalism need to be further validated. As the data are collected by the same method, common method variance should be considered. Due to the existence of social desirability, self-reported materialist values affected by social and culture cognition. Third, the cultural impact is very complicated. Future research can further explore the impact of culture on materialistic values by means of experimental approaches.

The results of the present studies have some implications for marketing. For example, China is shifting from a planned economy to a market economy. Chinese consumers are experiencing dramatic changes within the country’s economic system and consumer markets (Xiao, 1997). It is necessary to educate Chinese consumers on how to avoid economic, physical, and psychological damage and irrational behaviors caused by the market trap (Jin et al., 2015). Similarly, a developing market in a country with a specific culture should also consider local cultural and behavioral characteristics and adapt localized marketing strategies (Sun et al., 2014). Specifically, since Chinese people closely relate happiness and success, emphasizing face (Liao and Wang, 2009) and addressing how products can bring both happiness and success could result in better outcomes.

AUTHOR CONTRIBUTIONS

JL: conceived the research idea; made the research design; performed the studies; analyzed the data; interpreted the results; prepared manuscript; LW: conceived the research idea; organized the research; provided test materials and equipment.

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REFERENCES

- Argyle, M., Martin, M., and Crossland, J. (1989). "Happiness 'as a function of personality and social encounters," in *Recent Advances in Social Psychology: An International Perspective*, eds J. P. Forgas and J. M. Innes (Amsterdam: Elsevier), 189–203.
- Barger, S. D., Donoho, C. J., and Wayment, H. A. (2009). The relative contributions of race/ethnicity, socioeconomic status, health, and social relationships to life satisfaction in the United States. *Qual. Life Res.* 18, 179–189. doi: 10.1007/s11136-008-9426-2
- Bauer, W. (1976). *China and the Search for Happiness: Recurring Themes in Four Thousand Years of Chinese Cultural History*, trans. M. Shaw. New York, NY: Seabury Press.
- Baumgartner, H., and Homburg, C. (1996). Applications of structural equation modeling in marketing and consumer research: a review. *Int. J. Res. Mark.* 13, 139–161. doi: 10.1016/0167-8116(95)00038-0
- Belk, R. W. (1984). Three scales to measure constructs related to materialism: reliability, validity, and relationships to measures of happiness. *Adv. Consum. Res.* 11, 291–297.
- Belk, R. W. (1985). Materialism: trait aspects of living in the material world. *J. Consum. Res.* 12, 265–280. doi: 10.1086/208515
- Belk, R. W., and Pollay, R. W. (1985a). Images of ourselves: the good life in twentieth century advertising. *J. Consum. Res.* 11, 887–897. doi: 10.1086/209024
- Belk, R. W., and Pollay, R. W. (1985b). "Materialism and magazine advertising during the twentieth century," in *Advances in Consumer Research*, Vol. 12, eds E. Hirschman and M. Holbrook (Provo, UT: Association for Consumer Research), 394–398.
- Bradley, B. P., Mogg, K., Falla, S. J., and Hamilton, L. R. (1998). Attentional bias for threatening facial expressions in anxiety: manipulation of stimulus duration. *Cogn. Emot.* 12, 737–753. doi: 10.1080/026999398379411
- Brown, K. W., Kasser, T., Ryan, R. M., and Konow, J. (2016). Materialism, spending, and affect: an event-sampling study of marketplace behavior and its affective costs. *J. Happiness Stud.* 17, 2277–2292. doi: 10.1007/s10902-015-9694-9
- Burroughs, J. E., and Rindfleisch, A. (2002). Materialism and well-being: a conflicting values perspective. *J. Consum. Res.* 29, 348–370. doi: 10.1086/344429
- Campbell, A., Converse, P. E., and Rodgers, W. L. (1976). *The Quality of American Life*. New York, NY: Russell Sage Foundation.
- Campbell, D. T. (1970). "Various social attitude scales," in *Measures of Social Psychological Attitudes*, eds J. P. Robinson and P. R. Shaver (Ann Arbor, MI: Institute for Social Research), 648–654.
- Chiasson, N., Dube, L., and Blondin, J. (1996). Happiness: a look into the folk psychology of four cultural groups. *J. Cross Cult. Psychol.* 27, 673–691. doi: 10.1177/0022022196276002
- Devellis, R. (1991). *Scale Development: Theory and Application*. Newbury Park, CA: Sage publications, 1991.
- Dittmar, H., Bond, R., Hurst, M., and Kasser, T. (2014). The relationship between materialism and personal well-being: a meta-analysis. *J. Pers. Soc. Psychol.* 107, 879–924. doi: 10.1037/a0037409
- Dolan, P., Peasgood, T., and White, M. (2007). Do we really know what makes us happy? A review of the economic literature on the factors associated with subjective well-being. *J. Econ. Psychol.* 29, 94–122. doi: 10.1016/j.joep.2007.09.001
- Eberhardt, J. L., Goff, P. A., Purdie, V. J., and Davies, P. G. (2004). Seeing black: race, crime, and visual processing. *J. Pers. Soc. Psychol.* 87, 876–893. doi: 10.1037/0022-3514.87.6.876
- Fan, C., and Karnilowicz, W. (1997). Measurement of definitions of success among Chinese and Australian girls. *J. Cross Cult. Psychol.* 28, 598–599. doi: 10.1177/0022022197285005
- Feather, N. T. (1998). Attitudes toward high achievers, self-esteem, and value priorities for Australian, American, and Canadian students. *J. Cross Cult. Psychol.* 29, 749–760. doi: 10.1177/0022022198296005
- Furnham, A., and Cheng, H. (2000a). Lay theories of happiness. *J. Happiness Stud.* 1, 227–246. doi: 10.1023/A:1010027611587
- Furnham, A., and Cheng, H. (2000b). Perceived parental behavior, self-esteem and happiness. *Soc. Psychiatr. Epidemiol.* 35, 463–470.
- Ger, G., and Belk, R. W. (1990). Measuring and comparing materialism cross-culturally. *Adv. Consum. Res.* 17, 186–192.
- Ger, G., and Belk, R. W. (1996). Cross-cultural differences in materialism. *J. Econ. Psychol.* 17, 55–77. doi: 10.1016/0167-4870(95)00035-6
- Gibbons, F. X., and Buunk, B. P. (1999). Individual differences in social comparison: development of a scale of social comparison orientation. *J. Pers. Soc. Psychol.* 76, 129–142. doi: 10.1037/0022-3514.76.1.129
- Gotlib, I. H., McLachlan, A. L., and Katz, A. N. (1988). Biases in visual attention in depressed and nondepressed individuals. *Cogn. Emot.* 2, 185–200. doi: 10.1080/02699938808410923
- Heaney, J. G., Goldsmith, R. E., and Jusoh, W. J. W. (2005). Status consumption among Malaysian consumers: exploring its relationships with materialism and attention-to-social-comparison-information. *J. Int. Consum. Mark.* 17, 83–98. doi: 10.1300/J046v17n04_05
- Hills, P., and Argyle, M. (2002). The Oxford Happiness Questionnaire: a compact scale for the measurement of psychological well-being. *Pers. Individ. Di.* 33, 1073–1082. doi: 10.1016/S0191-8869(01)00213-6
- Himmelfarb, S., and Lickteig, C. (1982). Social desirability and the randomised response technique. *J. Pers. Soc. Psychol.* 43, 710–717. doi: 10.1037/0022-3514.43.4.710
- Ho, D. Y. F. (1976). On the concept of face. *Am. J. Sociol.* 81, 867–884. doi: 10.1086/226145
- Hofstede, G. (1980). *Cultures Consequences*. London: Sage Publications.
- Hong, L., Cao, H., and Fang, G. (2009). Development of money attitude scale and research on its validity and reliability. *Chin. J. Clin. Psychol.* 17, 139–141.
- Hu, H. C. (1944). The Chinese concept of "face". *Am. Anthropol.* 46, 45–64. doi: 10.1525/aa.1944.46.1.02a00040
- Jin, X. T., Wang, H. F., Wang, T. X., Li, Y., and Deng, S. L. (2015). Why Chinese elites buy what they buy the signalling value of conspicuous consumption in China. *Int. J. Market Res.* 57, 877–908. doi: 10.2501/IJMR-2015-041
- Kasser, T., Ryan, R. M., Zax, M., and Sameroff, A. J. (1995). The relations of maternal and social environments to late adolescents' materialistic and pro-social values. *Dev. Psychol.* 31, 907–914. doi: 10.1037/0012-1649.31.6.907
- Kilbourne, W., Grünhagen, M., and Foley, J. (2005). A cross-cultural examination of the relationship between materialism and individual values. *J. Econ. Psychol.* 26, 624–641. doi: 10.1016/j.joep.2004.12.009
- Krebs, D. L. (1982). "Altruism-A rational approach," in *The Development of Prosocial Behavior*, ed. N. Eisenberg (New York, NY: Academic Press), 53–76. doi: 10.1016/B978-0-12-234980-5.50008-0
- Li, Y. A., and Chen, Y. L. (2013). Dimensional structure of oxford happiness questionnaire (Revision) and verification of its reliability and validity. *Health Med. Res. Pract.* 10, 34–41.
- Liao, J. Q., and Wang, L. (2009). Face as a mediator of the relationship between material value and brand consciousness. *Psychol. Mark.* 26, 987–1001. doi: 10.1002/mar.20309
- Lu, L. (2001). Understanding happiness: a look into the Chinese folk psychology. *J. Happiness Stud.* 2, 407–432. doi: 10.1023/A:1013944228205
- Lu, L., and Shih, J. B. (1997). Sources of happiness: a qualitative approach. *J. Soc. Psychol.* 137, 181–187. doi: 10.1080/00224549709595429
- Macleod, C., Mathews, A., and Tata, P. (1986). Attentional bias in emotional disorders. *J. Abnorm. Psychol.* 95, 15–20. doi: 10.1037/0021-843X.95.1.15
- Marcus, H. R., and Kitayama, S. (1991). Culture and self: implications for cognition, emotion, and motivation. *Psychol. Rev.* 98, 224–253. doi: 10.1037/0033-295X.98.2.224
- Marcus, H. R., and Kitayama, S. (1994). A collective fear of the collective: Implications for selves and theories of selves. *Pers. Soc. Psychol. Bull.* 20, 568–579. doi: 10.1177/0146167294205013
- Mason, R. S. (1981). *Conspicuous Consumption: A Study of Exceptional Consumer Behavior*. Westmead: Gower.
- Mogg, K., and Bradley, B. P. (1999). Some methodological issues in assessing attentional biases for threatening faces in anxiety: a replication study using a modified version of the probe detection task. *Behav. Res. Ther.* 37, 595–604. doi: 10.1016/S0005-7967(98)00158-2
- Moschis, G. P., and Churchill, G. A. (1978). Consumer socialization: a theoretical and empirical analysis. *J. Mark. Res.* 15, 599–609. doi: 10.2307/3150629
- Mullen, M. R. (1995). Diagnosing measurement equivalence in cross-national research. *J. Int. Bus. Stud.* 3, 573–596. doi: 10.1057/palgrave.jibs.8490187
- Nagpaul, T., and Pang, J. S. (2017). Materialism lowers well-being: the mediating role of the need for autonomy - correlational and experimental evidence. *Asian J. Soc. Psychol.* 20, 11–21. doi: 10.1111/ajsp.12159

- Nisbett, R. E., Peng, K., Choi, I., and Norenzayan, A. (2001). Culture and systems of thought: holistic versus analytic cognition. *Psychol. Rev.* 108, 291–310. doi: 10.1037/0033-295X.108.2.291
- Ogden, H. J., and Venkat, R. (2001). Social comparison and possessions: Japan vs. Canada. *Asia Pac. J. Mark. Logist.* 13, 72–84. doi: 10.1108/13555850110764775
- Otero-Lopez, J. M., Pol, E. V., Bolano, C. C., and Marino, M. J. S. (2011). Materialism, life-satisfaction and addictive buying: examining the causal relationships. *Pers. Individ. Di.* 50, 772–776. doi: 10.1016/j.paid.2010.12.027
- Pandelaere, M. (2016). Materialism and well-being: the role of consumption. *Curr. Opin. Psychol.* 10, 33–38. doi: 10.1016/j.copsyc.2015.10.027
- Paulhus, D. L. (1984). Two-component models of socially desirable responding. *J. Pers. Soc. Psychol.* 46, 598–609. doi: 10.1037/0022-3514.46.3.598
- Richins, M. L. (1987). A multivariate analysis of responses to dissatisfaction. *J. Acad. Mark. Sci.* 15, 24–31. doi: 10.1007/BF02722168
- Richins, M. L. (1994). Special possessions and the expression of material values. *J. Consum. Res.* 21, 522–533. doi: 10.1086/209415
- Richins, M. L. (2004). The material values scale: measurement properties and development of a short form. *J. Consum. Res.* 31, 209–219. doi: 10.1086/383436
- Richins, M. L. (2013). When wanting is better than having: materialism, transformation expectations, and product-evoked emotions in the purchase process. *J. Consum. Res.* 40, 1–18. doi: 10.1086/669256
- Richins, M. L., and Chaplin, L. N. (2015). Material parenting: how the use of goods in parenting fosters materialism in the next generation. *J. Consum. Res.* 41, 1333–1357. doi: 10.1086/680087
- Richins, M. L., and Dawson, S. (1992). A consumer values orientation for materialism and its measurement: scale development and validation. *J. Consum. Res.* 19, 303–316. doi: 10.1086/209304
- Ruiz-Caballero, J. A., and Bermudez, J. (1997). Anxiety and attention: is there an attentional bias for positive emotional stimuli? *J. Gen. Psychol.* 124, 194–210. doi: 10.1080/00221309709595517
- Sampson, E. E. (1988). The debate on individualism: indigenous psychologies of individual and their role in personal and societal functioning. *Am. Psychol.* 43, 15–22. doi: 10.1037/0003-066X.43.1.15
- Schlenker, B. R., and Weingold, M. F. (1989). “Goals and the self-identification process: constructing desires and identities,” in *Goal Concepts in Personality and Social Psychology*, ed. L. A. Pervin (Hillsdale, NJ: Lawrence Erlbaum Associates), 243–290.
- Schneider, W., Eschman, A., and Zuccolotto, A. (2002). *E-Prime Users Guide*. Pittsburgh, PA: Psychology Software Tools Inc.
- Segerstrom, S. C. (2001). Optimism and attentional bias for negative and positive stimuli. *Pers. Soc. Psychol. Bull.* 27, 1334–1343. doi: 10.1177/01461672012710009
- Shi, J. Q., Wang, L., and Ddeng, W. (2005). Reliability and validity of the subjective well-being scale of Chinese middle school students. *Chin. Mental Health J.* 19, 727–729.
- Singhal, R., and Misra, G. (1992). Meaning of achievement: the role of socio-cultural background and social class. *Psychologia* 35, 42–49.
- Sun, G., D'Alessandro, S., and Johnson, L. (2014). Traditional culture, political ideologies, materialism and luxury consumption in China. *Int. J. Consum. Stud.* 38, 578–585. doi: 10.1111/ijcs.12117
- Tang, T. L. P. (1993). The meaning of money: extension and exploration of the money ethic scale in a sample of university students in Taiwan. *J. Organ. Behav.* 14, 93–99. doi: 10.1002/job.4030140109
- Tang, T. L. P. (1995). The development of a short money ethic scale: attitudes toward money and pay satisfaction revisited. *Pers. Individ. Di.* 19, 809–816. doi: 10.1016/S0191-8869(95)00133-6
- Tang, T. L. P., Luna-Arocas, R., and Whiteside, H. D. (2003). Money ethic endorsement, self-reported income, and life satisfaction: University faculty in the US and Spain. *Pers. Rev.* 32, 756–773. doi: 10.1108/00483480310498701
- Tatzel, M. (2002). “Money worlds” and well-being: an integration of money dispositions, materialism and price-related behavior. *J. Econ. Psychol.* 23, 103–126. doi: 10.1016/S0167-4870(01)00069-1
- Ting-Toomey, S. (1988). “Intercultural conflict styles: a face negotiation theory,” in *Theories in Intercultural Communication*, eds Y. Y. Kim and W. B. Gudykunst (Newbury Park, CA: Sage), 213–235.
- Ting-Toomey, S. (1998). Facework competence in intercultural conflict: an updated face-negotiation theory. *Int. J. Intercult. Relatsh.* 22, 187–225. doi: 10.1016/S0147-1767(98)00004-2
- Triandis, H. C. (1994). “Theoretical and methodological approaches to the study of collectivism and individualism,” in *Individualism and Collectivism: Theory, Method and Applications*, eds U. Kim, H. C. Triandis, C. Kagitcibasi, S. Choi, and G. Yoon (Thousand Oaks, CA: Sage), 41–51.
- Triandis, H. C. (1995). *Individualism and Collectivism*. Boulder, CO: Westview.
- Triandis, H. C., McCusker, C., and Hui, C. H. (1990). Multimethod probes of individualism and collectivism. *J. Pers. Soc. Psychol.* 59, 1006–1020. doi: 10.1037/0022-3514.59.5.1006
- Wang, H. L. (2005). *The Development and Influence Factors of Money Ethics Attitude*. Master dissertation, Peking University, Beijing.
- Wang, M. J., Wang, L., and Shi, J. Q. (2006). Reliability and validity of the social comparison orientation scale. *Chin. Ment. Health J.* 20, 302–305.
- Webster, C., and Beatty, R. C. (1997). Nationality, materialism, and possession importance. *Adv. Consum. Res.* 24, 204–210.
- West, S. G., Finch, J. F., and Curran, P. J. (1995). “Structural equation models with nonnormal variables: problems and remedies,” in *Structural Equation Modeling: Concepts, Issues, and Applications*, ed. R. H. Hoyle (Thousand Oaks, CA: Sage), 56–75.
- Wheaton, B. (1987). Assessment of fit in over-identified models with latent variables. *Sociol. Methods Res.* 16, 118–154. doi: 10.1177/0049124187016001005
- Wong, N. Y., and Ahuvia, A. C. (1998). Personal tastes and family face: self-concepts and luxury consumption in confucian and Western societies. *Psychol. Mark.* 15, 423–441. doi: 10.1002/(SICI)1520-6793(199808)15:5<423::AID-MAR2>3.0.CO;2-9
- Xiao, J. J. (1997). “Chinese consumer movement,” in *Encyclopedia of the Consumer Movement*, eds S. Brobeck, R. N. Mayer, and R. O. Herrmann (Santa Barbara, CA: ABC-CLIO), 104–109.
- Yamauchi, K., and Templer, D. (1982). The development of a money attitudes scale. *J. Pers. Assess.* 46, 522–528. doi: 10.1207/s15327752jpa4605_14

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